

**PhD in Economics**  
**1st year Econometrics test (2018-07-09 — resit)**

Name: \_\_\_\_\_

1. Say if the following statements are unambiguously true (TRUE), unambiguously false (FALSE) or impossible to classify the way they are stated (NOT NECESSARILY). Write the motivations to your answers **only** in the space provided. A “Not necessarily” answer with no motivations will be considered wrong.

- (a) If  $A$  and  $B$  are positive definite square matrices of the same dimension, then  $A - B$  is also positive definite.

True  False  Not necessarily

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- (b) If  $X$  and  $Y$  are two random variables and  $E(X) = 0$ , then  $\text{Cov}(X, Y) = E(X \cdot Y)$ .

True  False  Not necessarily

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- (c) If  $X_n \xrightarrow{p} 0.5$ , then  $\ln(X_n) - \ln(1 - X_n) \xrightarrow{p} 1$ .

True  False  Not necessarily

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- (d) If  $\sqrt{n}(X_n - 0.5) \xrightarrow{d} N(0, 0.25)$ , then  $\sqrt{n}[\ln(X_n) - \ln(1 - X_n)] \xrightarrow{d} N(0, 4)$ .

True  False  Not necessarily

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- (e) You have a sample of  $n$  iid realisations of a random variable  $X$ , with  $E(X) = \mu$ ; consider the statistic  $\hat{\mu} = \sqrt{\bar{X} \cdot V}$ , where  $\bar{X}$  is the sample mean of  $x_i$  and  $V$  is the sample variance. Then,  $\hat{\mu}$  is a consistent estimator of  $\mu$ .

True  False  Not necessarily

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2. You are reading a paper about the impact of attending (at least) a training course in the last 12 months on the probability of being employed. The author of the paper estimated a linear probability model (LPM) where the dependent variable is a dummy indicator equal to 1 if the individual is employed at the moment of the interview and 0 otherwise (*emp*). The regressor of primary interest is also a dummy indicator, equal to 1 if the individual attended (at least) a training course in the 12 months before the interview and 0 otherwise (*tr*). The estimated equation is therefore

$$emp_i = \delta \cdot tr_i + \mathbf{x}_i\boldsymbol{\beta} + u_i, \quad (1)$$

where  $u_i$  is the error term and  $\mathbf{x}_i$  are further individual characteristics.

Table 1 reports descriptive statistics of the dependent variable, the training indicator and further individual characteristics used in the analysis.

Table 1: Summary statistics of the variables used in the analysis

Variable	Mean	Std. Dev.	Minimum	Maximum
Employment status	0.6921	0.4616	0	1
Training indicator	0.0666	0.2493	0	1
Female	0.5309	0.4991	0	1
Age	43.498	10.047	26	64
Primary education	0.0872	0.2821	0	1
Secondary education	0.2084	0.4062	0	1
Tertiary education	0.7045	0.4563	0	1
Presence of kids younger than 12	0.3539	0.4782	0	1
Observation			33,348	

The estimation results of Equation (1) are reported in Table 2.

Table 2: Estimation results of employment equation

	Coefficient		Standard error <sup>(a)</sup>
Training indicator	0.1793	***	0.0055
Female	-0.2044	***	0.0044
Age	0.0644	***	0.0018
Age squared	-0.0009	***	0.0000
<i>Education - Reference: Primary education</i>			
Secondary education	-0.0868	***	0.0086
Tertiary education	-0.0587	***	0.0075
Presence of kids younger than 12	-0.1207	***	0.0051
Constant	-0.0923	**	0.0399
Observations			33,348
$R^2$			0.2118
$F(7, 33340)$			1,772.26

Notes: \* Significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.

<sup>(a)</sup> Standard errors are robust to heteroskedasticity.

After reading Table 2 answer the following questions:

- What is the marginal effect of attending (at least) a training course on the probability of being employed?
- Compute the t-statistic for the test of significance of the estimated parameter of the training indicator.
- Why does the author of this paper computed standard errors robust to heteroskedasticity? Would you have done it as well? Motivate your answer.
- What is the marginal effect of one more year of age on the probability of being employed?

- (e) Why is the the indicator for primary education omitted from the set of regressors included in the model specification?

The paper also presents a second estimated employment equation. The equation is similar to the previous one, with the only difference that there is one more regressor, which is the interaction between the training indicator and the female indicator. The estimation results are presented in Table 3.

Table 3: Estimation results of employment equation with gender heterogenous effect of training

	Coefficient		Standard error
Training indicator	0.0935	***	0.0063
Training indicator*Female	0.1841	***	0.0103
Female	-0.2167	***	0.0047
Age	0.0641	***	0.0018
Age squared	-0.0009	***	0.0000
<i>Education - Reference: Primary education</i>			
Secondary education	-0.0851	***	0.0086
Tertiary education	-0.0570	***	0.0075
Presence of kids younger than 12	-0.1190	***	0.0051
Constant	-0.0807	**	0.0400
Observations			33,348
$R^2$			0.2142
$F(8, 33339)$			1,596.84

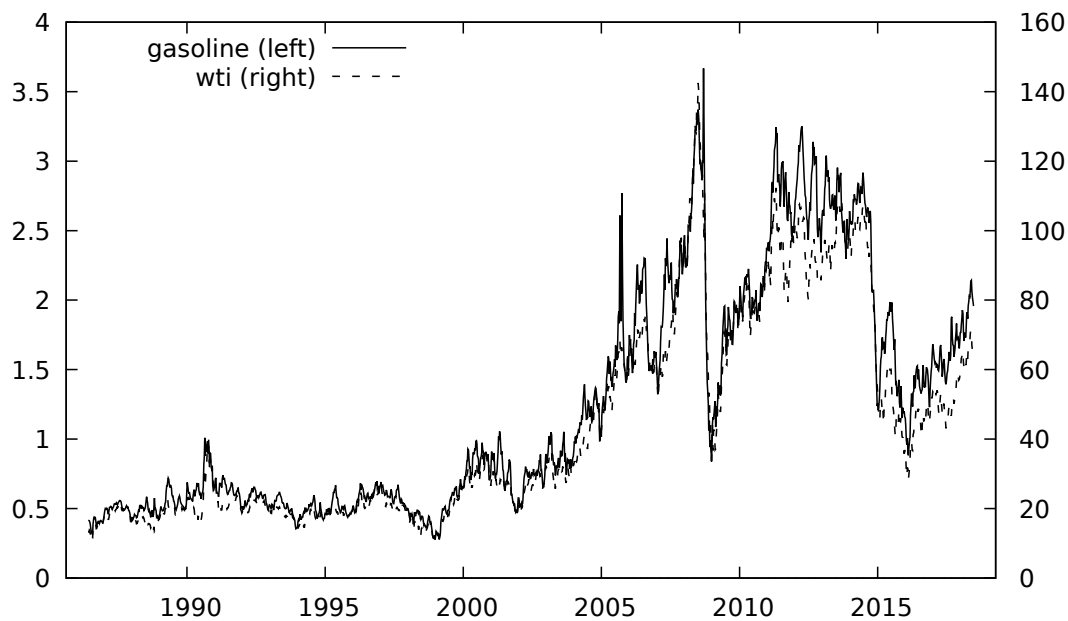
Notes: \* Significant at 10%; \*\* significant at 5%; \*\*\* significant at 1%.

(a) Standard errors are robust to heteroskedasticity.

After reading Table 3 answer the following questions:

- (f) Is the effect of training significantly different between males and females?  
 (g) What is the marginal effect of training on the probability of employment for men?  
 And for women?

Figure 1: Oil and gasoline prices



3. Figure 1 shows three weekly time series, from 06/06/1985 to 22/06/2018, defined as follows:

<b>Name</b>	<b>Definition</b>
gas	Natural logarithm of the spot price for U.S. Gulf Coast, regular gasoline (Source: eia.gov)
oil	Natural logarithm of the spot price for West Texas Intermediate crude oil price (Source: eia.gov)
markup	gas - oil

Economic theory would suggest that, since crude oil is by far the main component of the cost of producing gasoline, the long-run elasticity of the price of gasoline with respect to the price of oil should be 1.

Annex 1 contains a gretl script performing various tests and estimating procedures on these data, together with the unedited output from the program. Write **in no more than one page** the main conclusions you are able to draw from the output (and please please please, write legibly); if you don't comply with the space limitation, you will get 0 points, even if what you write is worth the Nobel Prize.

# Annex 1

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## Gretl script file

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```
# static regression
ols gas const oil

# take differences
d_oil = diff(oil)
d_gas = diff(gas)
# ecm model
scalar p = 6
ols d_gas const d_gas(-1 to -p) d_oil(0 to -p) gas(-1) oil(-1)
# Godfrey test
modtest 13 -a --quiet

# unit root tests
adf 13 oil --c --test-down=BIC
adf 13 gas --c --test-down=BIC
adf 13 markup --c --test-down=BIC

# lag selection
var 13 oil gas --lagselect
p = iminc($test[,2]) # use BIC

# Johansen test
coint2 p gas oil
# VECM
vecm p 1 gas oil

restrict
    b1 + b2 = 0
end restrict
```

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## Output

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Model 1: OLS, using observations 1986-06-06:2018-06-22 (T = 1673)

Dependent variable: gas

	coefficient	std. error	t-ratio	p-value	
const	-3.53383	0.0130809	-270.2	0.0000	***
oil	0.989228	0.00361230	273.8	0.0000	***
Mean dependent var	-0.011093	S.D. dependent var	0.657400		
Sum squared resid	15.74986	S.E. of regression	0.097085		
R-squared	0.978204	Adjusted R-squared	0.978191		
F(1, 1671)	74993.80	P-value(F)	0.000000		
Log-likelihood	1528.842	Akaike criterion	-3053.684		
Schwarz criterion	-3042.839	Hannan-Quinn	-3049.666		
rho	0.911448	Durbin-Watson	0.176628		

Model 2: OLS, using observations 1986-07-11:2018-06-22 (T = 1668)

Dependent variable: d\_gas

	coefficient	std. error	t-ratio	p-value	
const	-0.335603	0.0377083	-8.900	1.43e-18	***
d_gas_1	0.146425	0.0245458	5.965	2.98e-09	***
d_gas_2	-0.0733443	0.0248388	-2.953	0.0032	***
d_gas_3	0.103174	0.0245671	4.200	2.82e-05	***
d_gas_4	0.0514291	0.0246190	2.089	0.0369	**
d_oil	0.754124	0.0226904	33.24	5.28e-186	***
d_oil_1	-0.0205386	0.0292229	-0.7028	0.4823	
d_oil_2	0.0688646	0.0293081	2.350	0.0189	**
d_oil_3	-0.0378244	0.0288432	-1.311	0.1899	
d_oil_4	-0.0990733	0.0288769	-3.431	0.0006	***
gas_1	-0.0954441	0.0105602	-9.038	4.33e-19	***
oil_1	0.0939659	0.0105511	8.906	1.36e-18	***

Mean dependent var	0.000989	S.D. dependent var	0.050424		
Sum squared resid	2.375611	S.E. of regression	0.037875		
R-squared	0.439511	Adjusted R-squared	0.435788		
F(11, 1656)	118.0511	P-value(F)	4.3e-199		
Log-likelihood	3099.351	Akaike criterion	-6174.703		
Schwarz criterion	-6109.670	Hannan-Quinn	-6150.605		
rho	0.001530	Durbin-Watson	1.996874		

Excluding the constant, p-value was highest for variable 12 (d\_oil\_1)

Breusch-Godfrey test for autocorrelation up to order 13

Test statistic: LMF = 1.361811,  
with p-value =  $P(F(13,1643) > 1.36181) = 0.171$

Alternative statistic:  $TR^2 = 17.781324$ ,  
with p-value =  $P(\text{Chi-square}(13) > 17.7813) = 0.166$

Ljung-Box  $Q' = 12.9229$ ,  
with p-value =  $P(\text{Chi-square}(13) > 12.9229) = 0.454$

Augmented Dickey-Fuller test for oil  
testing down from 13 lags, criterion BIC  
sample size 1669

unit-root null hypothesis:  $a = 1$

test with constant  
including 3 lags of (1-L)oil  
model:  $(1-L)y = b_0 + (a-1)y(-1) + \dots + e$   
estimated value of  $(a - 1)$ : -0.00267861  
test statistic:  $\tau_c(1) = -1.7302$   
asymptotic p-value 0.416  
1st-order autocorrelation coeff. for e: -0.003  
lagged differences:  $F(3, 1664) = 17.063$  [0.0000]

Augmented Dickey-Fuller test for gas  
testing down from 13 lags, criterion BIC  
sample size 1669  
unit-root null hypothesis:  $a = 1$

test with constant  
including 3 lags of (1-L)gas  
model:  $(1-L)y = b_0 + (a-1)y(-1) + \dots + e$   
estimated value of  $(a - 1)$ : -0.00393731  
test statistic:  $\tau_c(1) = -2.13623$   
asymptotic p-value 0.2305  
1st-order autocorrelation coeff. for e: 0.005  
lagged differences:  $F(3, 1664) = 25.501$  [0.0000]

Augmented Dickey-Fuller test for markup  
testing down from 13 lags, criterion BIC  
sample size 1668  
unit-root null hypothesis:  $a = 1$

test with constant  
including 4 lags of (1-L)markup  
model:  $(1-L)y = b_0 + (a-1)y(-1) + \dots + e$   
estimated value of  $(a - 1)$ : -0.103058  
test statistic:  $\tau_c(1) = -9.49483$   
asymptotic p-value 1.349e-17  
1st-order autocorrelation coeff. for e: 0.003  
lagged differences:  $F(4, 1662) = 15.387$  [0.0000]

VAR system, maximum lag order 13

The asterisks below indicate the best (that is, minimized) values of the respective information criteria, AIC = Akaike criterion, BIC = Schwarz Bayesian criterion and HQC = Hannan-Quinn criterion.

lags	loglik	p(LR)	AIC	BIC	HQC
1	5978.33476		-7.195584	-7.176013	-7.188330
2	6014.30278	0.00000	-7.234100	-7.201482	-7.222010
3	6030.09295	0.00000	-7.248305	-7.202640	-7.231379
4	6046.36271	0.00000	-7.263088	-7.204375*	-7.241326
5	6056.61223	0.00040	-7.270617	-7.198858	-7.244020*
6	6058.40761	0.46421	-7.267961	-7.183154	-7.236528
7	6060.97515	0.27372	-7.266235	-7.168381	-7.229966
8	6064.49229	0.13409	-7.265653	-7.154752	-7.224548
9	6073.26383	0.00152	-7.271402	-7.147454	-7.225461
10	6077.79432	0.05959	-7.272041*	-7.135046	-7.221264
11	6078.28398	0.91292	-7.267812	-7.117770	-7.212199
12	6083.75084	0.02732	-7.269579	-7.106490	-7.209131
13	6084.44908	0.84481	-7.265601	-7.089465	-7.200317

Johansen test:

Number of equations = 2  
Lag order = 4  
Estimation period: 1986-07-04 - 2018-06-22 (T = 1669)  
Case 3: Unrestricted constant

Log-likelihood = 10793.3 (including constant term: 6056.9)

Rank	Eigenvalue	Trace test	p-value	Lmax test	p-value
0	0.047752	84.795	[0.0000]	81.664	[0.0000]
1	0.0018739	3.1305	[0.0768]	3.1305	[0.0768]

Corrected for sample size (df = 1660)

Rank	Trace test	p-value
0	84.795	[0.0000]
1	3.1305	[0.0767]

eigenvalue      0.047752      0.0018739

beta (cointegrating vectors)

gas	-11.082	-0.010940
oil	10.967	1.5362

alpha (adjustment vectors)

gas	0.0064122	-0.0017113
oil	-0.0023742	-0.0017179

renormalized beta

gas	1.0000	-0.0071209
oil	-0.98959	1.0000

renormalized alpha

gas	-0.071061	-0.0026290
oil	0.026311	-0.0026391

long-run matrix (alpha \* beta')

	gas	oil
gas	-0.071043	0.067693
oil	0.026330	-0.028677

VECM system, lag order 4

Maximum likelihood estimates, observations 1986-07-04-2018-06-22 (T = 1669)

Cointegration rank = 1

Case 3: Unrestricted constant

beta (cointegrating vectors, standard errors in parentheses)

gas	1.0000 (0.0000)
oil	-0.98959 (0.015082)

alpha (adjustment vectors)

gas	-0.071061
oil	0.026311

Log-likelihood = 6055.337

Determinant of covariance matrix = 2.4193787e-06

AIC = -7.2347

BIC = -7.1762

HQC = -7.2130

Equation 1: d\_gas



	coefficient	std. error	t-ratio	p-value	
const	-0.250569	0.0469438	-5.338	1.07e-07	***
d_gas_1	0.176205	0.0316676	5.564	3.06e-08	***
d_gas_2	-0.0199701	0.0315120	-0.6337	0.5263	
d_gas_3	0.122651	0.0315490	3.888	0.0001	***
d_oil_1	0.0531366	0.0375437	1.415	0.1572	
d_oil_2	-0.0247674	0.0368505	-0.6721	0.5016	
d_oil_3	0.00690781	0.0371647	0.1859	0.8526	
EC1	-0.0710615	0.0132732	-5.354	9.82e-08	***

Mean dependent var 0.000953 S.D. dependent var 0.050430  
Sum squared resid 3.976672 S.E. of regression 0.048930  
R-squared 0.062571 Adjusted R-squared 0.058620  
rho 0.004186 Durbin-Watson 1.991006

Equation 2: d\_oil

	coefficient	std. error	t-ratio	p-value	
const	0.0938062	0.0395033	2.375	0.0177	**
d_gas_1	0.0419184	0.0266483	1.573	0.1159	
d_gas_2	0.0853092	0.0265174	3.217	0.0013	***
d_gas_3	0.0310184	0.0265485	1.168	0.2428	
d_oil_1	0.100191	0.0315931	3.171	0.0015	***
d_oil_2	-0.148400	0.0310097	-4.786	1.86e-06	***
d_oil_3	0.0633167	0.0312741	2.025	0.0431	**
EC1	0.0263114	0.0111694	2.356	0.0186	**

Mean dependent var 0.000951 S.D. dependent var 0.042077  
Sum squared resid 2.815980 S.E. of regression 0.041175  
R-squared 0.046449 Adjusted R-squared 0.042430  
rho 0.000711 Durbin-Watson 1.995420

Cross-equation covariance matrix:

	gas	oil
gas	0.0023827	0.0012652
oil	0.0012652	0.0016872

determinant = 2.41938e-06

Restriction:

$$b[1] + b[2] = 0$$

Test of restrictions on cointegrating relations

eigenvalue 1 = 0.0474912

Unrestricted loglikelihood (lu) = 6055.337

Restricted loglikelihood (lr) = 6055.1083

2 \* (lu - lr) = 0.457361

P(Chi-square(1) > 0.457361) = 0.49886